



The Labor Market Impact of Immigration: What Have We Learned?

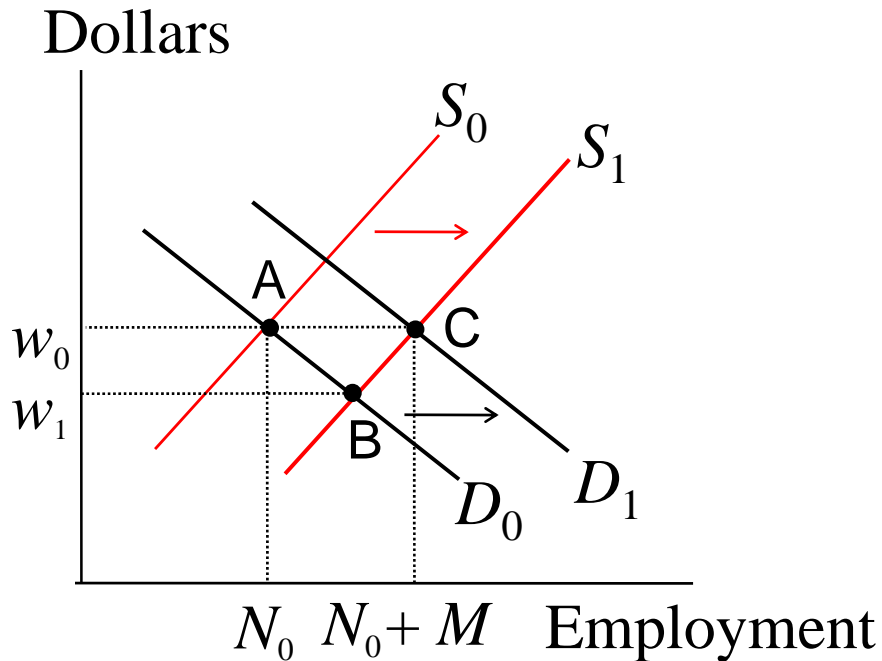
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1. The question

- Do immigrants alter the employment opportunities of native workers?
 - “After World War I, laws were passed severely limiting immigration. Only a trickle of immigrants has been admitted since then. . .By keeping labor supply down, immigration policy tends to keep wages high. Let us underline this basic principle: Limitation of the supply of any grade of labor relative to all other productive factors can be expected to raise its wage rate; an increase in supply will, other things being equal, tend to depress wage rates”
 - Paul Samuelson, *Economics*, 1964.

2. Short and long run, homogeneous labor



- Pre-immigration equilibrium at A.
- **Short-run:** stock of capital fixed. Equilibrium at B. Wages and native employment fall.
- **Long-run:** returns to capital fixed. Equilibrium at C. Wages and native employment return to pre-immigration equilibrium if the aggregate production function $f(K, L)$ is linearly homogeneous.



3. Short and long run, heterogeneous labor

- Production function: $q = f(K, L_1, \dots, L_n)$.
- **Short run:**
 - Wage of the “average” labor input falls.
 - There are also distributional impacts: drop in relative wage of groups with largest immigrant-induced supply shifts.
- **Long run:**
 - Wage change for “average” labor input is zero.
 - But there are distributional impacts: reduction in relative wage of groups with largest immigrant-induced supply shifts.



4. The empirical puzzle

- Summary of findings (*circa* 2000):
 - Friedberg and Hunt (1995, p. 42): “The effect of immigration on the labor market outcomes of natives is small.”
 - National Academy of Sciences (1997, p. 220): “The weight of the empirical evidence suggests that the impact of immigration on the wages of competing native workers is small.”
- Puzzle: The immigrant supply shock in the United States has been very large, and Hamermesh (1993) concludes that the labor demand curve is not perfectly elastic. **Why can't we observe an impact?**



5. Spatial correlation approach

- Suppose there are a number of *closed* labor markets that immigrants penetrate *randomly*. We can then relate the change in the wage in a particular market to the number of immigrants in that market. The estimated correlation measures the impact of immigration.
- Immigrants cluster in a small number of geographic areas. Most studies exploit this geographic clustering to test the implications of the textbook model.
 - Grossman (1982), Borjas (1987), Altonji and Card (1991), LaLonde and Topel (1991), Schoeni (1997), Card (2001). Most influential study: Card's (1990) analysis of the Mariel flow.



6. Card's (1990) Mariel paper

Unemployment rate of blacks in:	The Mariel flow (Card, 1990)		The Mariel flow that didn't happen (Angrist & Krueger, 1999)	
	Before (1979)	After (1981)	Before (1993)	After (1995)
Miami	8.3	9.6	10.1	13.7
Comparison cities	10.3	12.6	11.5	8.8

The comparison cities are Atlanta, Houston, Los Angeles, and Tampa-St. Petersburg.



7. Problems with spatial correlations


- Immigrants may not be randomly distributed across labor markets. If immigrants cluster in cities with thriving economies, there would be a spurious positive correlation between immigration and local employment conditions (Borjas, 2001).
- Local labor markets are not closed. Natives may respond to the immigrant supply shock by moving their labor or capital to other cities, thereby re-equilibrating the national economy.
 - There is an unresolved debate over whether these equilibrating flows exist. See Borjas, Freeman, Katz (1997), Card (2001), Borjas (2006).
- Measurement error (Aydemir and Borjas, 2006)



8. A tangential puzzle

- Explicit contradiction between spatial correlations estimated in the minimum wage and immigration literatures.
- Card-Krueger et al estimate regressions of the form:
$$\log L = \alpha \log w + \text{other variables}$$
 - Use regions as the unit of observation.
 - Estimated coefficient α is often close to zero.
- Grossman, Card, et al estimate regressions of the form:
$$\log w = \beta \log L + \text{other variables}$$
 - Use regions as the unit of observation
 - Estimated coefficient β is often close to zero.
- Both sets of results cannot be correct. Labor demand curve cannot be perfectly elastic and perfectly inelastic at the same time!
- Interesting conjecture: Neither conclusion is right.

9. A new type of natural experiment



INCREASED WAGES

**HIRING IMMEDIATELY
GENERAL PRODUCTION WORKERS
DAY & EVENING SHIFTS**
(Limited Van Service Available
to and from Crider)

**Please apply in person
Monday - Thursday
8:00 A.M. - 10:00 A.M.**

**At
1 Plant Ave., Hwy 57 Stillmore, GA
EOE**

- “After a wave of raids by federal immigration agents on Labor Day weekend, a local chicken-processing company called Crider Inc. lost 75% of its mostly Hispanic 900-member work force. The crackdown threatened to cripple the economic anchor of this fading rural town. But for local African-Americans, the dramatic appearance of federal agents presented an unexpected opportunity. Crider suddenly raised pay at the plant. An advertisement in the weekly Forest-Blade newspaper blared “Increased Wages” at Crider, starting at \$7 to \$9 an hour—more than a dollar above what the company had paid many immigrant workers.” (*The Wall Street Journal*, January 17, 2007)
- Implied wage elasticity = -0.20



10. Factor proportions approach

- Borjas, Freeman, and Katz (1992, 1997) change the unit of analysis to the national level. If the aggregate technology is given by a linear homogeneous CES production function with two inputs (high school dropouts and everyone else), the relative wage of the two groups depends linearly on their relative quantities.

$$\log (w_1/w_2) = \text{constant} + \sigma \log(L_1/L_2)$$

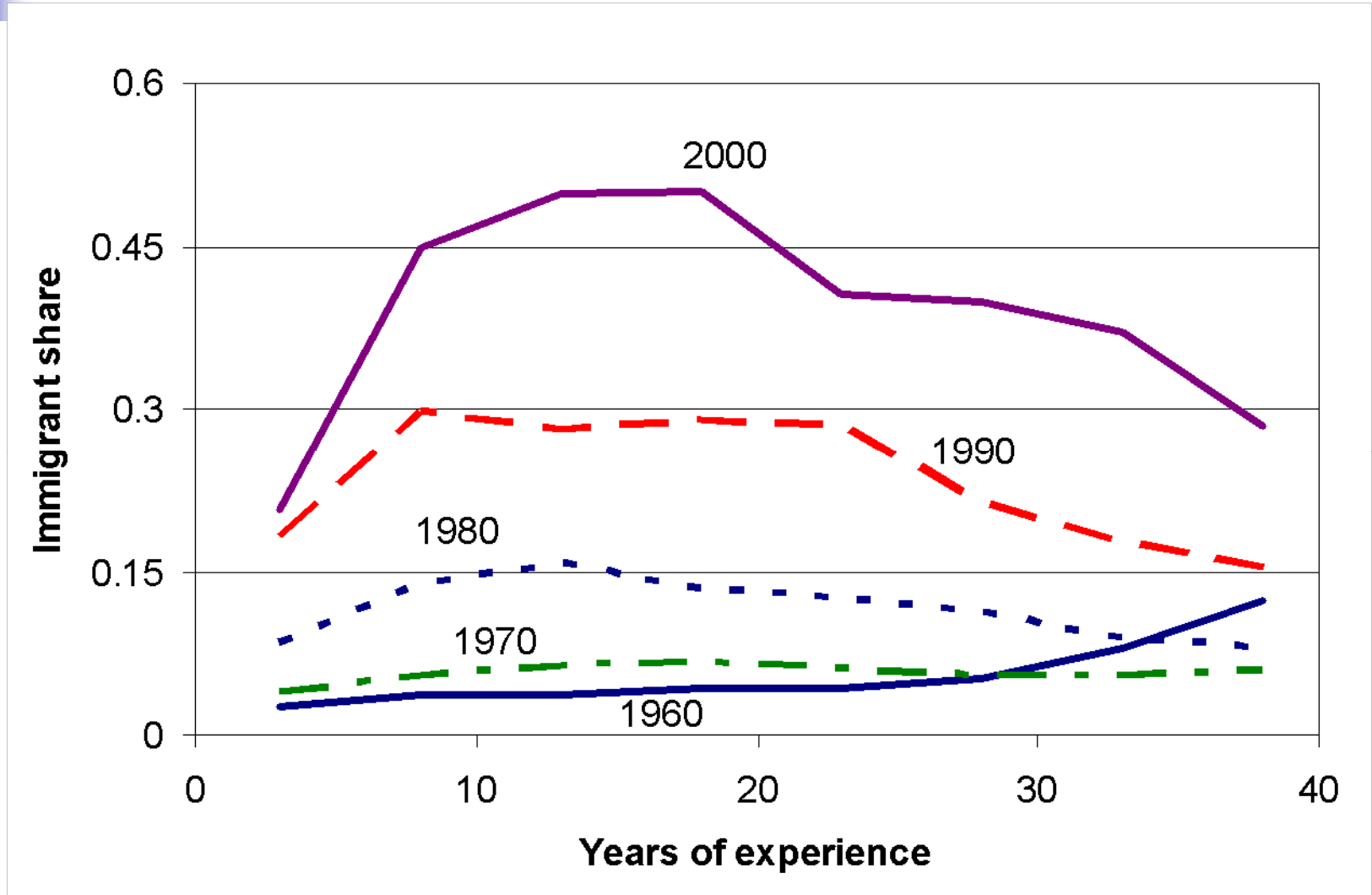
- BFK use Katz-Murphy estimate of σ to predict the value of the wage ratio between skilled and unskilled natives if immigration had not changed factor proportions. They find that 40% of the decline in the relative wage of high school dropouts is due to immigration.
- **Problem:** The factor proportions approach does not *estimate* the impact of immigration on the labor market. Instead, it *simulates* the impact.



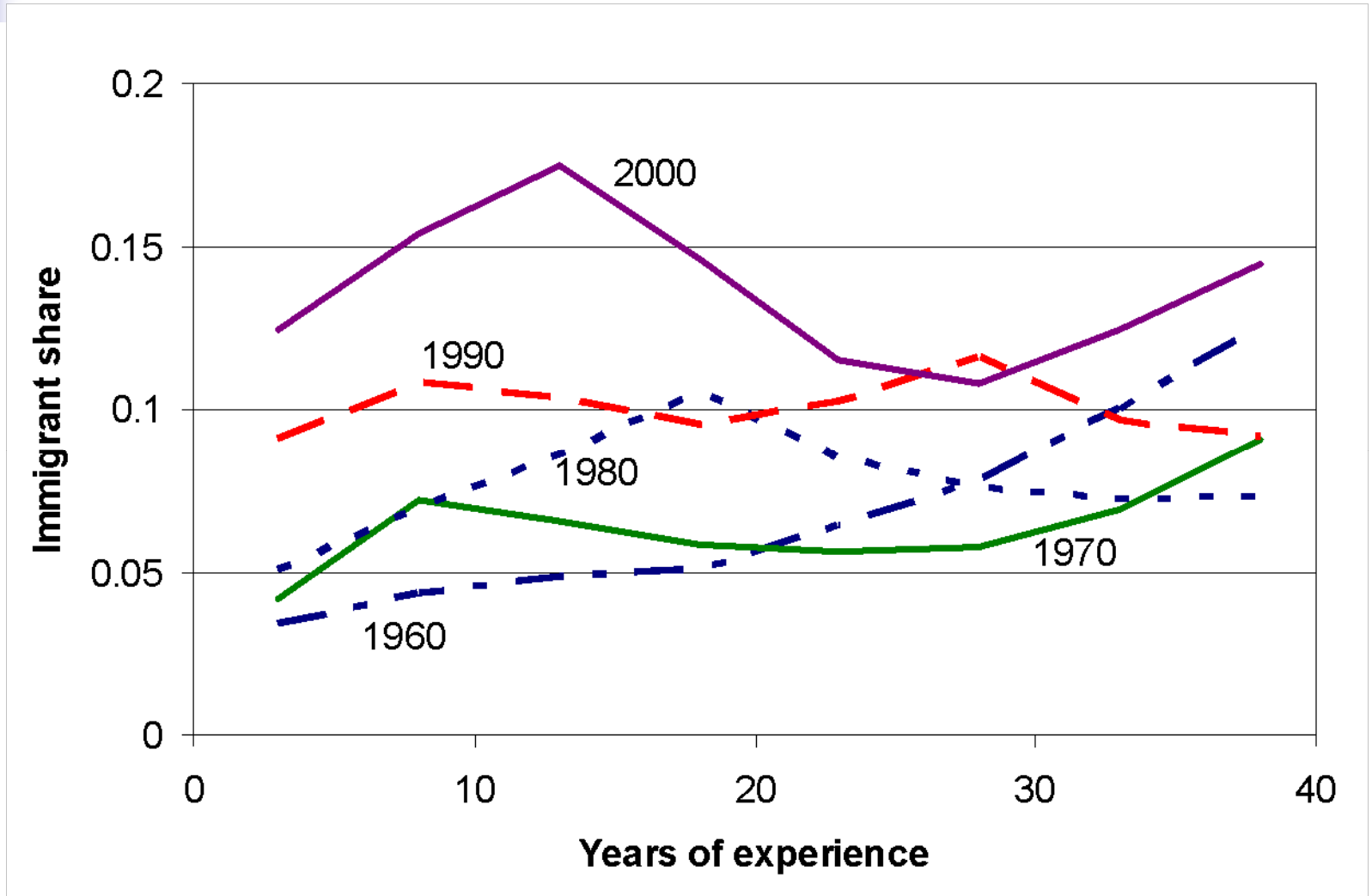
11. An alternative approach (Borjas, QJE, 2003)

- First, pay closer attention to the definition of a skill group. Both schooling *and* work experience determine a person's stock of acquired skills.
 - Immigration is not balanced evenly across all experience cells in a particular schooling group. The immigrant influx will tend to affect some native workers more than others. And the nature of the supply "imbalance" changes over time.
- Second, local labor market may not be the right unit of observation. So examine evolution of national wage structure (in the spirit of Murphy-Welch, 1992; Katz-Murphy, 1992, Card and Lemieux, 2001).
- This approach reconfirms that the labor demand curve is indeed downward sloping: An influx of immigrants into a particular skill group lowers the wage of that skill group.

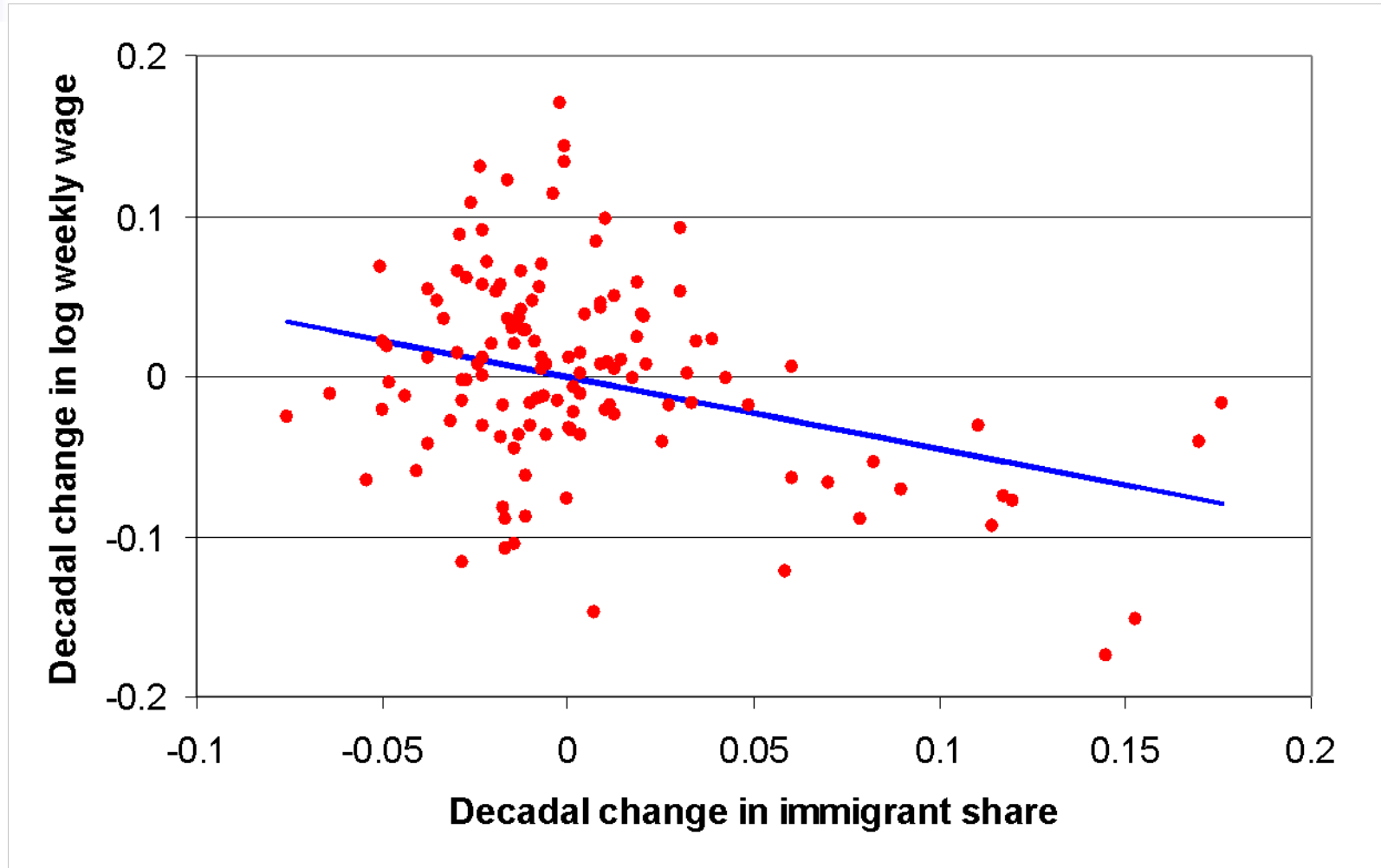
12. Supply shock for high school dropouts in US



13. Supply shock for college graduates in US



14. Scatter diagram relating wages and immigration (removing decade effects)





15. Regression model

- Let y_{ijt} be the mean value of a particular labor market outcome for *native* men with education s , experience x , in year t . I stack the data across skill groups and calendar years and estimate:

$$y_{sxt} = \theta p_{sxt} + S + X + T + (S \times X) + (S \times T) + (X \times T) + \varepsilon_{sxt}$$

- S are fixed effects indicating educational attainment; X are fixed effects indicating work experience; T are fixed effects indicating calendar year; p is the immigrant share in the cell.
- The interactions control for the experience profile of y differing across schooling groups, and for the impact of education and experience changing over time.



16. Interpreting the adjustment coefficient

- The regression model is:

$$y_{sxt} = \theta p_{sxt} + S + X + T + (S \times X) + (S \times T) + (X \times T) + \varepsilon_{sxt}$$

- Let $m_{sxt} = M_{sxt} / N_{sxt}$; the percentage increase in the labor supply of the skill group. Define the wage elasticity as:

$$\frac{\partial \log w_{sxt}}{\partial m_{sxt}} = \theta (1 - p_{sxt})^2.$$

- By 2000, the immigrant share in the workforce was around 0.14. The factor price elasticity—evaluated at the mean value of the immigrant share—is obtained by multiplying θ by 0.7.

17. Estimated adjustment coefficients

Study	Immigration	Emigration
1. United States (Aydemir and Borjas, 2007)	-0.489	---
	(0.223)	
2. Canada (Aydemir and Borjas, 2007)	-0.507	---
	(0.202)	
3. Mexico (Aydemir and Borjas, 2007)	---	0.798
		(0.443)
4. Mexico (Mishra, 2006)	---	0.440
		(0.110)
5. Great Britain (Peev, 2007)	-0.508	---
	(0.198)	
6. Puerto Rico (Borjas, 2007)	-0.583	0.405
	(0.267)	(0.184)



18. Remarks

- Does approach require the existence of a “national” labor market?
- Not necessarily. Suppose that local labor markets are isolated island economies. The national level regression (which averages out the local data) should then exactly replicate the spatial correlation.
- The fact that the two approaches lead to different answers seems to contradict the assumption that local labor markets are isolated island economies.



19. A structural approach

- A structural approach specifies the technology of the aggregate production function.
- One can then estimate the own-effect of immigrants on the wage of competing native workers and the cross-effects on the wage of other natives.
- Suppose the aggregate production function can be represented in terms of a three-level CES technology: Similarly educated workers with different levels of work experience are aggregated to form the effective supply of an education group; and workers across education groups are then aggregated to form the national workforce.



20. The three-level CES technology

$$Q_t = \left[\lambda_{Kt} K_t^\nu + \lambda_{Lt} L_t^\nu \right]^{1/\nu}, \text{ with } \nu = 1 - \frac{1}{\sigma_{KL}}.$$

$$L_t = \left[\sum_s \theta_{st} L_{st}^\rho \right]^{1/\rho}, \text{ with } \rho = 1 - \frac{1}{\sigma_E}.$$

$$L_{st} = \left[\sum_x \alpha_{sx} L_{sxt}^\eta \right]^{1/\eta}, \text{ with } \eta = 1 - \frac{1}{\sigma_X}.$$



21. Advantages of three-level CES

- There are 33 factors of production: 32 education-experience groups plus capital.
- A more general specification (e.g., the translog) requires estimating 561 different parameters. The three-level CES reduces the parameter space to three elasticities of substitution.
- **Main disadvantage:** the CES greatly restricts the types of substitution that can exist.
 - σ_X takes on the same value for workers in adjacent experience categories as for workers who differ greatly in their experience; σ_E is the same for high school dropouts and high school graduates as for high school dropouts and college graduates; and σ_{KL} is the same for all the different types of workers.
 - Results are extremely sensitive to assumption about the trend in the productivity vector θ .



22. Measuring the wage impact of immigration

- Use estimated σ 's to calculate the wage impact of the 1980-2000 flow of international migrants.
- Theory predicts $w = w(K, L_1, \dots, L_n)$. Suppose the capital stock is constant. Let $\varepsilon_{sx,ij}$ be the factor price elasticity (the wage impact on group sx of an increase in the number of workers in group ij). Let m_{ij} be the migrant-induced percent increase in labor supply.
- Short-run impact on the log wage of group (s, x) is:

$$\Delta \log w_{sx} = \sum_i \sum_j \varepsilon_{sx,ij} m_{ij}$$



23. Long-run impact

- Long run: suppose rental rate of capital is constant. The predicted wage impact is:

$$\Delta \log w_{sx} = \frac{s_K}{\sigma_{KL}} \tilde{K} + \sum_i \sum_j \varepsilon_{sx,ij} m_{ij}, \quad \text{where } \tilde{K} = \sum_{\ell} s_{\ell} m_{\ell} / s_L$$

- Short- and long-run impacts differ only by a positive constant, so full capital adjustment mutes the adverse wage impact of immigration but leaves the relative wage effects unchanged.
- Because first-level CES has constant returns, supply shocks induce a corresponding shift in the capital stock that leaves the capital/labor ratio constant and the average wage in the economy unchanged.

24. Predicted percent wage impacts of 1980-2000 immigrant supply shock, United States and Canada

	High school dropouts	High school grads.	Some college	College grads.	Post-grad. degree	All workers
U.S.						
Supply shock (%)	22.8	8.1	9.2	12.0	12.8	11.1
Short-run, $\sigma_{KL} = 1.0$	-7.1	-2.3	-2.7	-3.6	-3.9	-3.3
Long run	-3.8	1.0	0.6	-0.3	-0.6	0.0
Canada						
Supply shock (%)	-5.6	3.2	13.3	29.0	32.0	10.4
Short-run, $\sigma_{KL} = 1.0$	2.7	-0.9	-4.3	-9.7	-11.5	-3.3
Long-Run	6.0	2.4	-1.0	-6.4	-8.2	0.0

25. Predicted percent wage impacts of 1980-2000 emigrant supply shock, Mexico

	High school dropouts, 0-8 yrs	High school dropouts, 9-11 yrs	High school grads.	Some college	College grads.	All workers
Supply shock (%)	-11.6	-11.6	-35.6	-29.7	-5.8	-14.6
Short-run, $\sigma_{KL} = 1.0$	3.8	3.8	12.3	10.2	1.7	4.8
Long run	-1.0	-1.1	7.5	5.4	-3.1	0.0



26. Reflections on the long run

- This type of long-run modeling cannot be correct.
- In the long run, the average wage goes back to its pre-immigration equilibrium in both receiving and sending countries.
- What about factor price equalization?
- In absence of information about extent of capital adjustment, the CES approach is perhaps best interpreted as giving the impact of immigration on **relative wages**, rather than the impact of immigration on absolute wages.
- Important caveat: CES approach imposes a lot of restrictions that may not be valid.



27. Imperfect substitution (Ottaviano-Peri, 2006)

- Suppose immigrants and natives within a skill group are not perfect substitutes.

$$L_{sxt} = \left[\theta_{sxt} N_{sxt}^{\gamma} + (1 - \theta_{sxt}) M_{sxt}^{\gamma} \right]^{1/\gamma},$$

- Ottaviano-Peri claim this extension of the model leads to a new simulation result: 1990-2000 immigrant influx raises average wage of native workers by 1 to 2 percent
- **BUT:** average wage change must be zero, so average wage of immigrants must have declined substantially!
- **Another caveat:** Finding of imperfect substitution is not robust to specification of model. In US data, one cannot reject hypothesis of perfect subst. if one includes educ*time, experience*time, and educ*experience fixed effects.

28. Predicted long run wage impact of 1990-2004 inflow with imperfect substitution

First 2 columns from O-P, Table 9	Natives	Immigrants	Observed 1990-2000 change in $\Delta \log w$ (M-N)
HS dropouts	-1.1%	-16.3%	+0.002
College graduates	+0.7%	-24.2%	-0.005

- Impact of immigrants on immigrants seems implausibly high. The raw data **contradicts** implication that relative wage of immigrants dropped dramatically in the 1990s.
- Aggregated results are similar to Borjas *QJE* results. Borjas defines “natives” as the pre-existing population. So wage of pre-existing HS dropout dropped by about -5.7%.
- One interpretation: CES restrictions are straight-jacketing the data and twisting the evidence.



29. Conclusion

- Wage response to immigration is a crucial parameter in any assessment of the efficiency and distributional impact of international migration.
- Previous estimates—particularly those at the MSA level—may have underestimated the impact because of attenuation bias introduced by sampling error and behavioral responses.
- A number of puzzles and confusing conjectures remain:
 - Why do spatial correlations lead to contradictory conclusions in the minimum wage and immigration literatures?
 - How do “pre-existing workers” respond to immigration?
 - How does capital respond to immigration?
 - How long does it take to reach the long run?
 - Does immigration lead to factor price equalization?